1 2	WHERE ARE THE ELECTRIC VEHICLES? A SPATIAL MODEL FOR VEHICLE-CHOICE COUNT DATA
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4	T. Donna Chen
5	The University of Texas at Austin
6	6.9E Cockrell Jr. Hall
7	Austin, TX 78712-1076
8	<u>donna.chen(a)utexas.edu</u>
9	
10	Yiyi Wang
11	Assistant Professor
12	Montana State University
13	Civil Engineering Department
14	y1y1.wang@ce.montana.edu
15	
16	Kara M. Kockelman
17	(Corresponding author)
18	Professor and William J. Murray Jr. Fellow
19	Department of Civil, Architectural and Environmental Engineering
20	The University of Texas at Austin
21	kkockelm(a)mail.utexas.edu
22	Phone: 512-471-0210
23	
24	
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30 21	Electric encliente (EV) and une distant to increase in marghest changes and the margin for the margin data
31	Electric vehicles (EVS) are predicted to increase in market share as auto manufacturers introduce
32 22	more rule efficient venicles to meet stricter CAFE standards and driver concerns of increasing
33	fuel costs. Reflecting spatial autocorrelation while controlling for a variety of demographic and
34	locational (e.g., built environment) attributes, this zone-level spatial count model in this paper
35	offers valuable information for power providers and charging station location decisions. By
36	anticipating over 745,000 personal-vehicle registrations across a sample of 1000 census block
37	groups in the Philadelphia region, a trivariate Poisson-lognormal conditional autoregressive
38	(CAR) model anticipates Prius hybrid EV, other EV, and conventional vehicle ownership
39	levels. Initial results signal higher EV ownership rates in more central zones with higher
40	household incomes, along with significant residual spatial autocorrelation, suggesting that
41	spatially-correlated latent variables and/or peer (neighbor) effects on purchase decisions are
42	present. Such data sets will become more comprehensive and informative as EV market shares
43	rise. This work's multivarate Poisson-lognormal CAR modeling approach offers a rigorous,
44	behaviorally-defensible framework for spatial patterns in choice behavior.
16	

1 **KEY WORDS:** Electric vehicles, spatial count modeling, vehicle choice, vehicle ownership,

- 2 consumer behavior, conditional auto-regressive, built environment
- 3

4 MOTIVATION

- 5 As auto manufacturers introduce a variety of new vehicles to meet stricter fuel economy
- 6 standards, in the U.S. and abroad, and driver concerns regarding long-term energy prices remain
- 7 high, hybrid electric vehicle (HEV), plug-in hybrid electric vehicle (PHEV), and battery electric
- 8 vehicle (BEV) sales are on the rise (Schweinberg 2013). After tracking the electric vehicle (EV)
- 9 market for 13 years, IDTechEX predicts that 35% of all vehicles in the world will be electric by
- 10 2025, with a likely mix of 25% hybrids and 10% BEVs (Harrop and Das 2012). With such
- 11 meaningful market share changes on the horizon, an ability to predict which households or
- neighborhoods are most likely to own such vehicles can provide important insights and opportunities for power-grid planning (e.g., updating transformers in certain locations).
- opportunities for power-grid planning (e.g., updating transformers in certain locations),
 transportation investments (e.g., identifying where public charging stations should be installed
- 14 transportation investments (e.g., identifying where public charging stations should be installed 15 for maximum utilization), and air quality policy-making (e.g., forecasting ozone level changes
- when more vehicle-miles are electrified).
- 17 Spatial patterns in growing EV ownership may also illuminate what causes or alleviates
- 18 owner/consumer "range anxiety" (i.e., a user's concern for being stranded with a fully discharged
- battery and no reasonable recharge option [Tate et al. 2008]). As illustrated by Khan and
- 20 Kockelman (2012), a 75-mile all-electric range (AER) BEV (like the 2013 Nissan LEAF) may
- be a very reasonable vehicle for 27% of single-vehicle households and nearly 70% of multiple-
- vehicle households in Seattle to own, assuming existing travel patterns, a willingness to charge the vehicle once a day, and a willingness to charge more than once or find another travel option
- 25 the venicle once a day, and a winnigness to charge more than once of find another travel option 24 (e.g., a relative's car or rental vehicle) just 3 days a year, on average (in order to exceed the
- 25 BEV's AER).
- 26 Recent evidence from the U.S. Department of Energy's and ECOtality's EV Project (Smart et al.
- 27 2013) suggests that 73% of miles driven by Americans in a Chevy Volt stay within its 35-mile
- 28 (EPA-rated) AER (thereby avoiding much gasoline use in this PHEV). However, according to
- 29 *Consumer Reports*' Car Brand Perception Survey (Bartlett 2012), range anxiety remains
- 30 consumers' top concern consumer regarding a possible EV purchase. Studies suggest that such 31 anxiety may fall as drivers become more familiar with EV technology and EV use (see, for
- anxiety may ran as drivers become more raminar with EV technology and EV use (see, for
 example, Wellings et al. [2011] and Taylor [2009]). As with open-road tolling, adaptive cruise
- 32 example, weilings et al. [2011] and Taylor [2009]). As with open-road toring, adaptive cruise 33 control, and other relatively new transport policies and technologies, it seems very possible that
- 34 potential owners will worry less about EV range limitations as they are exposed to EVs on local
- 35 roads, in neighbors' driveways, and nearby parking garages (Mau et al. 2008). Related to this,
- Axsen et al. (2009) surveyed over 1000 vehicle owners in Canada and California and found that
- 37 willingness-to-pay (WTP) for HEVs rose with higher (existing)) HEV market penetration rates.
- 38 This study econometrically models ownership rates of EVs and conventional vehicles across
- 39 Philadelphia neighborhoods, while allowing for such neighbor (spatial autocorrelation) effects; it
- 40 applies a new multivariate count model, with both spatially-lagged and (aspatial) cross-response
- 41 correlation opportunities.

42 BACKGROUND

1 Most EV forecasts are simply an aggregate market share, across a nation or region, with

- 2 estimates widely varying. For example, the U.S. National Research Council (2010) predicted 13
- 3 million EVs on U.S. roads by 2030 (4% of total fleet) in the most probable scenario and 40
- 4 million EVs (13.3% of total fleet) in the maximum practical case, while the U.S. Energy
- 5 Information Administration (EIA 2013) recently forecasted just 3% of all U.S. light duty vehicle
- 6 (LDV) *sales* will be EVs by 2040. Simulating consumer behavior under a business-as-usual
- 7 (BAU) model, Clement-Nyns et al. (2010) projected EVs to reach 30% of the Belgian passenger-
- 8 vehicle fleet by 2030. McKinsey's (2011) survey suggested that in three of the world's
- 9 "megacities" (New York City, Shanghai, and Paris), EVs may hit 16% of vehicle sales by 2015.
- 10 Within the U.S. northeast corridor, Pike Research (2011) projects that Washington, D.C. and
- 11 Delaware will have the highest annualized penetration rates of EVs by 2017, at 4.6 and 4.5
- 12 percent, respectively. Paul et al.'s (2011) microsimulation of U.S. household holdings forecasted
- 13 7.6% of the fleet to be HEVs and PHEVs by 2035 under BAU, and 13.1% under a feebate plus
- 14 doubled-gas price scenario, ceteris paribus. Examining both demand (for vehicles, batteries, and
- 15 gasoline) and supply constraints (on production), Neubauer et al. (2012) projected California's
- 16 PHEV + BEV population to hit 500,000 sometime between 2018 and 2020.
- 17 Many researchers have employed discrete choice models at a disaggregate (individual or
- 18 household) level to explore various vehicle ownership decisions. For example, Brownstone et al.
- 19 (1996) analyzed data from a stated preference survey on alternative-fuel vehicles and found that
- 20 two-vehicle households with children express a greater WTP for cleaner (emissions-reducing)
- 21 vehicles, as compared to childless households. Erdem et al. (2010) employed an ordered probit
- (OP) model to examine Turkish consumers' WTP for HEVs and found that higher-income
 females, with more education and concerns about global warming, are more likely to purchase
- HEVs. The relationship between income and vehicle preference tends to be complicated by
- 25 household size: Paul et al. (2011) found that households with higher household income per
- 26 member tend to prefer smaller vehicles, but larger households generally prefer larger vehicles
- 27 (for seating-capacity reasons).
- 28 Land use characteristics also influence ownership decisions. For example, higher residential
- 29 densities are associated with lower vehicle ownership and usage levels (e.g., Zhao and
- 30 Kockelman [2002] and Fang [2008]). Holding other household attributes (control variables)
- 31 constant, Brownstone and Golob (2009) predicted density reductions of 1,000 housing units per
- 32 square mile (or 1.56 units per acre) to associated with another 1,000 miles per year of vehicle-
- miles traveled and 65 more gallons of fuel consumed per household (with 20 gallons of this
- 34 difference accounted by choice of more fuel-efficient vehicles in higher-density settings). The
- 35 choice of higher fuel-economy vehicles may be largely attributable to lower light-duty truck¹
- 36 (LDT) ownership in such settings: Brownstone and Fang's (2009) Bayesian multivariate OP
- 37 model associates a 50 percent residential density increase with a modest but statistically
- 38 significant reduction on LDT ownership levels, and a 610-mile annual per-truck VMT decrease.
- 39 In the same study, demand for passenger car ownership was estimated to be inelastic with respect
- 40 to residential density (Brownstone and Fang 2009), but fuel economy can change significantly
- 41 within the car fleet, leading to EV purchases, rather than say large, luxury cars, and thereby offer
- 42 substantial energy savings. Using a multiple discrete-continuous extreme value (MCDEV)

¹ In the U.S. the light-duty truck definition includes cargo vans, minivans, sport-utility vehicles, and pickup trucks weighing less than 8,500 lbs loaded (i.e., the gross vehicle weight rating).

1 specification, Bhat et al. (2009) also found that smaller vehicle sizes are more prevalent in

2 neighborhoods high in both residential and commercial densities. Beyond simple density

3 measures, Potoglou et al. (2008) found vehicle ownership to depend somewhat on land use

4 diversity and transit proximity. Khan et al. (2012) also investigated the linkage between vehicle

5 ownership and a host of built environment factors, including network structure, bus stop density,

6 land use mix and jobs density, using a standard negative binomial model.

7 This research addresses a gap in our current understanding of EV ownership decisions by

8 examining the effects of demographic and land use characteristics at the neighborhood (i.e.,

9 Census block group) level, rather than at a regional level or individual/household level. In this

10 way, the work is able to quantify spatial autocorrelation or "neighbor effects" that can emerge in

11 the adoption of new technologies, and to predict adoption rates over space, without requiring

12 details on individuals.

13 DATA DESCRIPTION

14 The analysis relies on April 2012 vehicle registration data collected for the southeastern region

15 of Pennsylvania, around Philadelphia, by the Delaware Valley Regional Planning Commission

16 (DVRPC). The data set covers 2,980 census block groups containing 2,225,595 personal-vehicle

17 registrations. 18,674 of these registered vehicles are considered EVs, and most of those (i.e.,

18 13,421 or 72%) are HEVs, like the (original/non-plug-in) Toyota Prius and Honda Civic HEV.

19

20 The response variables used in this paper's trivariate model are the number of Prius EVs, non-

21 Prius EVs, and conventional (internal combustion engine) vehicles (ICEVs) in each Census

22 block group. Predictor variables (i.e., covariates) include block-group demographics, such as

23 population, resident worker², and job densities, plus household income shares (by income range),

24 and built-environment attributes (including distance to the central business district or CBD

25 [downtown Philadelphia] and centerline-mile densities of primary and secondary roads), as

- 26 discussed below.
- 27

28 The demographic details were originally provided at the traffic analysis zone (TAZ) level, by the

29 regional MPO, DVRPC. TransCAD's overlay function was used to impute such attributes at the

30 block-group level, in order to spatially match them to the vehicle ownership data. Table 1

31 provides summary statistics of all variables at the Census block-group level.

Table 1. Summary	v Statistics of	f Model Vai	riables at the	e Zone Level

	Mean	Median	Std. Dev.	Min	Max
Number (#) of Households (#HHs)	513.2	455.7	277.5	0	2.494E+03
Resident Worker Density (# per acre)	8.899	5.561	10.90	0	208.1
Emp. Density (jobs per acre)	9.651	2.727	51.03	1.500E-05	1124.8
Pop. Density	21.34	12.722	25.03	0	360.2

² The term "resident workers" refers to those who list themselves as employed (either full or part time) and reside in the zone. Such individuals typically work outside the zone in question.

(persons per acre)						
Income>\$35K Density						
(# of households with incomes	4.08	2.599	5.578	0	147.7	
over \$35,000 per year, per acre)						
Centerline Density (centerline						
miles of roadway per square	5.087E+04	1.521E+04	1.124E+05	0	1.920E+06	
mile of zone area)						
Euclidean Distance to CBD	12.28	0.111	0.706	0	10.67	
(mi.)	12.30 9.111		9.700	U	49.07	
HH0wrk Density [*] (# per acre)	2.874	1.269	4.276	0	92.52	
HH1wrk Density (# per acre)	3.216	1.748	4.772	0	122.4	
HH2wrk Density (# per acre)	1.868	1.214	2.257	0	38.96	
HH3wrk Density (# per acre)	0.428	0.262	0.514	0	5.449	
Response Variables (Counts)						
ICEVs	736.9	598.0	558.0	1	14,697	
Non-Prius EVs	1.749	1	2.333	0	19	
Prius EVs	4.474	2	5.773	0	37	

After merging the demographic and employment count information from the DVRPC with the
vehicle count layer (originally provided by the Delaware Valley Regional Planning Commission),
there were 2,909 Census block groups with both the response variables (vehicle counts by type)
and explanatory variables available.

6

7 According to planners at the DVRPC, \$35,000 is an important household income threshold that 8 is specially tracked, in part due to distinct travel patterns; thus, the density of these households 9 was tested as a covariate. Densities for most variables were used here because the vehicle-10 ownership rate per zone is scaled by the number of households per zone. In other words, the negative binomial process modeled here (described in detail below) benefits from an exposure or 11 12 size term, which here is the number of households (or population) per zone. (In theory, if the 13 number of households or persons doubles, one expects the number of owned vehicles to double, 14 ceteris paribus, in each of the three categories modeled. As is common in the practice of count-15 model estimation, this work allows for a non-unitary parameter on the natural logarithm of

16 households per zone, in order to provide more flexibility than forcing the model to exactly scale.)

17

18 Network connectivity may also have a bearing on vehicle ownership. For example, abundance of

19 roadway resources invites more vehicular travel, indirectly increasing household vehicle

20 ownership. And poorly-connected networks can stymie direct travel while increasing congestion,

21 possibly incentivizing the purchase of more fuel-efficient vehicles (including EVs). In this model,

22 network features are proxied by the centerline-mile density of primary and secondary roads³.

³ Primary and secondary roadway networks were obtained from the 2011 Census Tiger/Line archive for the state of Pennsylvania, at <u>http://catalog.data.gov/dataset/tiger-line-shapefile-2012-series-information-file-for-the-primary-and-secondary-roads-state-bas</u> Primary roads essentially refer to freeways and expressways and other divided, limited-access highways that are state-maintained or part of the interstate highway system. Secondary roads are major arterial streets, with multiple lanes of traffic in each direction, at-grade intersections, and possibly undivided medians.

1 Centerline density (for primary and secondary roads) was computed as the ratio between miles of

- 2 roadway and area of block groups (in square miles).
- 3 The population and jobs data were used in map format to define the regional CBD. Both
- 4 densities spike in the central Philadelphia area, corresponding to the rectangular area shown in
- 5 Figure 1. For block groups within this rectangular (CBD) area, the distance-to-CBD values were
- 6 set to zero; for block groups outside of this zone, Euclidean distances from each block group's
- 7 centroid were computed, to the centroid of the CBD rectangle.
- 8

9 METHODOLOGY

10

11 Zone-level vehicle-ownership counts, by type of vehicle, can be modeled as a spatial count

12 context with a multivariate response vector (of size 3×1 here), such as the model proposed and

13 estimated by Wang and Kockelman (2013). Wang and Kockelman's (2013) new specification

14 allows for multiple, simultaneous integer count responses, as well as spatial autocorrelation (due

15 to missing variables that trend in space), zonal heterogeneity, and spatially-lagged⁴ and aspatial

16 cross-correlations (across counts). A three-level response model is specified here, with the first

17 stage (for vehicle count values) expressed as a Poisson process:18

19
$$y_{ik} \sim \text{Poisson}(\lambda_{ik})$$
 (1)

20

where y_{ik} is the observed vehicle count by type (k = 1 denotes Prius EVs, 2 denotes non-Prius EVs, and 3 denotes ICEVs) for the i^{th} block-group polygon of the five-county Delaware Valley region. The (non-negative) average crash rates for each zone and vehicle type, λ_{ik} , are defined in the second stage, as follows:

25 26

27

$$\lambda_{ik} = E_i^{\alpha} \cdot \exp(x_i'\beta_k + \phi_k + u_i) \tag{2}$$

where E_i^{α} is the exposure or scaling term, of household population (per zone), as discussed earlier. The vector of all spatial random terms $\boldsymbol{\phi} = (\boldsymbol{\phi}'_1, \boldsymbol{\phi}'_2, \boldsymbol{\phi}'_3)'$ is multivariate normally distributed, where $\boldsymbol{\phi}_k$ is an *n* by 1 vector of spatial random effects or errors for ownership rates of the type *k* vehicle, such that:

32

33
$$\begin{pmatrix} \boldsymbol{\phi}_1 \\ \boldsymbol{\phi}_2 \\ \boldsymbol{\phi}_3 \end{pmatrix} \sim N \begin{pmatrix} \begin{pmatrix} \boldsymbol{\mu}_1 \\ \boldsymbol{\mu}_2 \\ \boldsymbol{\mu}_3 \end{pmatrix}, \begin{bmatrix} \Sigma_{11} & \Sigma_{12} & \Sigma_{13} \\ \Sigma'_{12} & \Sigma_{22} & \Sigma_{23} \\ \Sigma'_{13} & \Sigma'_{23} & \Sigma_{33} \end{bmatrix} \end{pmatrix}$$
 (3)

34

35 where the *n* by 1 vector $\boldsymbol{\mu}_k$ indicates the average ownership rate for response/vehicle type *k* (*k* = 1, 2, 3), and Σ_{kl} are *n* by *n* matrices describing the covariance structure between response types *k* and *l*.

⁴ "Spatially-lagged cross-correlation" captures the correlation across different vehicle types that occurs across space (mathematically expressed by the component $\eta_1 W$ in matrix A), while aspatial correlation describes the standard correlation (like that in an aspatial multivariate count model) across different response types, as formulated by the remaining component, $\eta_0 I$ in matrix A.

- 1 The three-level scheme developed here decomposes the massive covariance structure into a 2 series of conditional distributions that allow for spatial autocorrelation (measured by parameters ρ_1 , ρ_2 , and ρ_3), aspatial cross-correlations (gauged by parameters $\eta_{0,23}$, $\eta_{0,13}$, and $\eta_{0,12}$), and 3 spatially-lagged cross-correlations (expressed by parameters $\eta_{1,23}$, $\eta_{1,13}$, and $\eta_{1,12}$). Details about 4 5 model specification are provided in the Appendix. 6 7 This three-level multivariate conditional autoregressive (MCAR) model structure was thus used 8 to simultaneously predict counts of registered Toyota Prius EVs, non-Prius EVs (all EVs 9 excluding the Prius), and ICEVs for the Philadelphia region. The parameters were estimated using Bayesian Markov-chain Monte Carlo sampling techniques, coded in R and WinBUGS, as 10 11 described in Wang and Kockelman (2013). Due to computing limitations with larger sample 12 sizes (from non-standard posterior likelihoods [i.e., likelihood functions that do not follow a
- 13 known distributional form] associated with discrete count process and complex covariance
- 14 structures), Table 2's parameter estimates come from an n = 1,000 block-group sub-sample for
- 15 central Philadelphia, rather than the original 2,909 block groups that exist across the wider
- 16 region (and have complete data, as discussed earlier). Figure 1 shows the spatial distribution of
- 17 these vehicle counts (per household).





Figure 1. Locations of Vehicle Counts (by Type) per Household across Philadelphia Zones (Clockwise: Prius EVs, non-Prius EVs, and ICEVs)

1 **RESULTS**

2

3 As shown in Table 2, several control variables are associated with higher levels of vehicle

- 4 ownership, per household (in the block group) -- regardless of vehicle type. For example,
- 5 resident worker density is associated with more vehicles of each type, as expected, since most
- 6 American workers drive alone to work (McKenzie et al. 2010)). Elasticities⁵ suggest that for a
- 7 one-percentage point increase in resident-worker density (a measure of local population density),
- average increases of 3.4, 3.7, and 5.1 percentage points are expected in Prius EV, non-Prius EV,
 and ICEV ownership rates (vehicles per households) across zones, respectively. Zonal
- household counts (serving as the size measure in its logarithmic form) and employment density
- are also positively correlated with greater vehicle counts in a zone, with fairly strong statistical
- 12 and practical significance. Population density is not practically significant here, for EV
- 13 ownership levels, once household counts and resident-worker densities are accounted for,
- 14 perhaps due to this variable being inflated by household members who do not drive (such as
- 15 young children and the disabled or very elderly). However, this term exhibits a strong marginal
- 16 effect on ICEV ownership rates, with an elasticity estimate of +2.7 percent, suggesting the higher
- 17 likelihood of larger-vehicle ownership rates for households with such non-drivers (e.g.,
- 18 Brownstone and Fang [2009] estimated higher rates of pickup-truck ownership in California).
- 19 Indeed, as of model year (MY) 2012, the only EVs offering seating capacity above five were the

20 Toyota Highlander HEV and the Chevrolet Tahoe/GMC Yukon/Cadillac Escalade HEVs (with

- 21 the latter three vehicles all built on the same platform). In the current car market, larger families
- 22 who require more passenger and cargo capacities simply have fewer choices, when seeking an
- EV versus an ICEV.
- 24

25 Not surprisingly, higher shares of households with annual incomes above \$35,000 are associated 26 with higher vehicle ownership rates of all three types. In contrast, a higher density of lower-27 income households (those with less than \$35,000 annual income) is negatively associated with 28 EV ownership rates, though positively associated with ICEV ownership rates. The price 29 premiums on new EVs as compared to their ICEV counterparts range from \$2,655 to \$6,160, for 30 MY 2012 vehicles (Cunningham 2012) and no doubt pose a barrier for lower income families 31 when considering EVs (despite the value of EVs' longer term fuel-cost savings, as computed in 32 Khan and Kockelman [2012] and Tuttle and Kockelman [2012]). Low EV ownership rates 33 among lower income households may also be attributed to the limited selection of used EVs in 34 the market, as compared to used ICEVs. Interestingly, more ownership of non-Prius EVs is 35 expected (with a highly elastic average zone-based elasticity of +2.6%) when the density of 36 households having over \$35,000 annual income increases by one percent, versus ownership of 37 Prius EVs (with an elasticity of just +0.8%). This finding may reflect the Prius' relative modest 38 price among all EVs as well as its longer history on the market compared to other EVs (thus

- 39 making more Priuses available on the used vehicle market).
- 40
- 41 Controlling for all other variables, greater distance to the regional CBD is associated with much
- 42 lower ownership of both EV types (k = 1, 2) and higher ICEV ownership: A one-percent increase
- 43 in distance to CBD is accompanied by an average 2.3% increase in ICEV ownership rates as well

⁵ Here, elasticities represent the median value of the (MCMC-based) distribution of the averaged percentage increase in vehicle ownership rates (averaged across the 1000 zones in the sample) following a one-percent increase in each covariate (across all sampled zones), holding all other covariate values (in the sample) constant.

1 as a 26.2% and 20% decrease, respectively, in Prius EV, non-Prius EV ownership rates at the

- 2 block group level. These results are highly elastic, suggesting that the CBD distance variable is
- proxying for a variety of relevant attributes that change with distance but were not available as
- 4 covariates here, such as parking prices and parking space availability (which track land values, 5 mostly), transit access, and so forth. The fuel economy (Michalek et al. 2011) and vehicle size
- benefits of EVs, coupled with range anxiety for households living further from the regional
- center, may also be at play, in this ownership vs. distance relationship. The parameter
- coefficients of the ln(#HHs) variable (i.e., the natural logarithm of household counts per zone)
- 9 across the three models were estimated to be 0.75, 0.8, and 0.67 for Prius EV, non-Prius EV, and

10 ICEV counts, respectively, which suggests fewer vehicles per household in zones with more

households (which may be due to more population-dense zones having more households, even

- 12 though population density was also controlled for here [in a linear way]).
- 13

14 Centerline-mile density appears to have no influence on ownership rates, ceteris paribus. Zero-

- 15 worker households are negatively associated with vehicle ownership rates across all types,
- 16 presumably due to lower travel needs and lower incomes. Negative association is also found

17 between the density of households with 3+ workers and EV ownership rates (with elasticities of -

18 0.6% and - 2.6%, respectively) and a positive correlation with ICEV ownership rates (+3.0%

19 elasticity). Higher jobs densities come with substantially higher average vehicle counts across all

20 types: the effects are very elastic, with values of 2.5%, 3.0% and 6.9%, for Prius-EVs, non-Prius

- 21 EVs, and ICEVs, respectively.
- 22

23 Seemingly random variations in all three vehicle ownership rates exhibit spatially clustering, as

24 measured by the autocorrelation coefficients ρ_1 , ρ_2 and ρ_3 . By model construction, these spatial

- effects capture missing variables that trend in space, such as parking prices, congestion, and land
- 26 use intensity, which affect vehicle ownership decisions. The remarkably high (aspatial)
- 27 correlation between non-Prius EV and Prius EV error terms (described via the parameter η_{012} ,
- with mean estimate +0.813) reveals that higher Prius EV counts are likely to co-exist with higher

non-Prius EV counts, as expected. However, their spatial correlation (gauged by the parameter

30 η_{112}) is practically insignificant (i.e., neighboring zones' Prius-EV counts have no bearing on a 31 zone's non-Prius EV counts). In other words, the cross-correlations among different vehicle

types are local (within a zone), after controlling for spatial autocorrelation across zones (from

- 33 missing variables).
- 34

35 ICEV ownership rates exhibit a modest *aspatial* correlation with Prius and non-Prius EV

- 36 ownership rates, as reflected by the parameters $\eta_{013} = +0.334$ (pseudo t-stat. = 1.689) and $\eta_{023} =$
- +0.269 (pseudo t-stat. = 1.407), respectively. In other words, higher ICEV ownership rates tend

38 to rise with EV ownership rates, which may come from underlying factors shared across all

39 ownership rate decisions, and not captured by the model's covariates (e.g., other demographic

- 40 factors and parking availability).
- 41
- 42 Interestingly, ICEVs show a rather weak spatially-lagged cross-correlation with the two EV
- 43 categories ($\eta_{113} = +0.073$ and $\eta_{123} = +0.058$). Using measures of spatial correlation, the results in
- 44 this paper support previous findings (using non-spatial models and arguments) that vehicle
- 45 technology adoption is often influenced via a "neighbor effect," where new technologies become

1 more desirable as market penetration increases nearby (see, e.g. Mau et al. [2008] and Axen et al.

2 3 [2009]).

1 80	ne 2. Paran	ieler Estii	nates for	the Zone-L	Level veni	cie Registra		
		Mean	Std. Dev.	Pseudo t-stat	2.5%	Median	97.5%	Elasticit
	1 (PriusEV)	-2.138	0.425	-5.037	-2.895	-2.105	-1.333	-
Constant	2 (Non- PriusEV)	-2.033	0.209	-9.709	-2.423	-2.030	-1.625	-
	3 (ICEV)	2.628	0.235	11.188	2.277	2.571	3.153	0.030
Resi. Worker	1	0.447	0.062	7.263	0.331	0.437	0.587	0.034
Density	2	0.303	0.058	5.236	0.202	0.323	0.382	0.037
(# per acre)	3	0.136	0.056	2.438	0.010	0.136	0.222	0.051
	1	0.753	0.069	10.963	0.616	0.751	0.888	0.031
ln(#HHs)	2	0.801	0.038	21.214	0.729	0.796	0.874	0.019
	3	0.668	0.038	17.797	0.579	0.674	0.728	0.027
Emp.	1	0.020	0.007	3.080	0.007	0.020	0.033	0.025
Density.	2	0.017	0.006	2.876	0.006	0.017	0.029	0.030
(# per acre)	3	0.003	0.003	1.049	-0.003	0.003	0.008	0.069
	1	-0.027	0.043	-0.624	-0.143	-0.017	0.035	-0.004
Pop. Density	2	-0.058	0.018	-3.168	-0.095	-0.054	-0.028	0.004
(# per acre)	3	0.010	0.015	0.666	-0.017	0.009	0.040	0.027
HH > \$35K	1	1.615	0.146	11.054	1.383	1.592	1.864	0.008
Density	2	1.446	0.148	9.770	1.191	1.476	1.694	0.026
(# per acre)	3	0.194	0.084	2.303	0.052	0.204	0.327	0.027
HH < \$35K	1	-1.294	0.257	-5.035	-3.114	-1.294	-0.121	-0.008
Density	2	-0.725	0.143	-5.065	-2.235	7.122	-5.188	-0.026
(# per acre)	3	1.239	0.063	19.61	2.214	1.239	0.128	0.027
Centerline	1	-0.726	0.522	-1.390	-1.775	-0.726	0.275	0.008
Density	2	-0.712	0.484	-1.472	-1.668	-0.708	0.231	-0.026
(1/mi.)	3	-0.093	0.291	-0.319	-0.666	-0.088	0.447	-0.321
	1	-0.029	0.008	-3.734	-0.044	-0.029	-0.014	-0.262
Dist. CBD	2	-0.029	0.007	-3.856	-0.041	-0.030	-0.012	-0.200
(mi)	3	0.003	0.005	0.556	-0.007	0.003	0.012	0.023
HH0wrk	1	-0.713	0.110	-6.507	-0.902	-0.724	-0.485	0.005
Density	2	-0.575	0.091	-6.294	-0.748	-0.572	-0.397	-0.026
(# per acre)	3	-0.241	0.051	-4.752	-0.341	-0.241	-0.143	0.026
HH1wrk	1	-1.473	0.087	-16.883	-1.659	-1.460	-1.308	0.008
Density	2	-1.194	0.095	-12.504	-1.361	-1.203	-0.997	-0.026
(# per acre)	3	-0.306	0.096	-3.178	-0.487	-0.295	-0.155	-0.023
HH2wrk	1	-2.458	0.142	-17.347	-2.738	-2.445	-2.180	-0.005

 Table 2. Parameter Estimates for the Zone-Level Vehicle Registration Model

Density	2	-1.809	0.156	-11.633	-2.121	-1.800	-1.526	-0.027
(# per acre)	3	-0.492	0.161	-3.051	-0.851	-0.464	-0.242	-0.025
HH>3wrk	1	-3.241	0.369	-8.776	-3.905	-3.228	-2.504	-0.006
Density	2	-2.571	0.421	-6.103	-3.456	-2.505	-1.772	-0.026
(# per acre)	3	-0.323	0.353	-0.916	-0.966	-0.283	0.302	0.030
η_{012}	2	0.813	0.056	14.524	0.694	0.815	0.911	-
η_{013}	;	0.334	0.198	1.689	0.124	0.252	0.745	-
η_{023}	\$	0.269	0.191	1.407	0.076	0.194	0.702	-
η_{112}	2	0.074	0.025	2.932	0.034	0.071	0.132	-
η_{113}		0.073	0.025	2.919	0.032	0.070	0.129	-
η_{123}		0.058	0.072	0.810	0.012	0.026	0.268	-
ρ_1		0.655	0.108	6.078	0.429	0.662	0.844	-
ρ ₂		0.491	0.102	4.831	0.293	0.492	0.687	-
ρ ₃		0.913	0.147	6.217	0.471	0.974	0.992	-
τ_{v1}		4.982	0.753	6.616	3.640	4.931	6.566	-
τ_{v2}		5.628	0.784	7.178	4.234	5.591	7.284	-
τ _{v3}		2.179	3.701	0.589	0.115	0.303	11.230	-
τ ₁		3.311	0.608	5.445	2.232	3.272	4.568	-
τ ₂		1.056	0.228	4.632	0.704	1.022	1.604	-
τ ₃		1.070	1.758	0.609	0.059	0.170	5.609	-

Note: Bolded elasticities are practically significant, suggesting more than 1% change in the expected vehicle ownership level, following a 1% increase in the associated covariate.

CONCLUSIONS

Using a trivariate Poisson-lognormal CAR model, this study forecast registered-vehicle counts
 across southeastern Pennsylvania to examine the effect of zone-level characteristics on vehicle

ownership rates (per resident household) of Prius EVs, non-Prius EVs, and ICEVs while

9 reflecting spatial autocorrelation. Results reveal that all three vehicle ownership rates rise with

household counts (per zone) and resident-worker densities. In particular, EV ownership is

predicted to rise with as households get closer to the CBD, perhaps due to the relatively small

size of most EVs, the diminished range anxiety issues (thanks to shorter trips), and desire for

13 greater fuel economy (in congested central-city driving conditions). The spatial clustering effects

identified in this study also point to the existence of missing variables that trend in space and

15 possibly a "neighbor effect" (from nearby households owning such vehicles). As a technology

becomes more prevalent, familiarity and perceived value of such technologies and products tends

to rise. The spatial autocorrelation exhibited by the results here suggest that EV purchase rates

18 will rise, via such familiarity, thanks to rising market penetration.

19

20 The results of this study underscore the high continuing demand for ICEVs, thanks in part to the

21 far larger choice set of vehicles in that class, versus the mostly smaller vehicles offered in the EV

classes to date. (ICEs presently offer greater passenger and cargo capacity.) In general, lower

23 income households are less likely to purchase EVs; but, when they do, they are more likely to

24 purchase Toyota Priuses (the EV model with the longest market history) than non-Prius EVs,

1 suggesting that vehicle affordability and familiarity, along with used-vehicle availability, may

- 2 also be influencing EV purchase decisions.
- 3

4 A household's decision to purchase an EV is also influenced by many relatively complex

5 residential and transportation characteristics, not included in this study's set of covariates. These

6 include such things as availability of designated parking for a home charger to be installed,

7 nearby public charging infrastructure, and commute distances (versus AER, in the case of BEVs).

8 Nevertheless, the model presented here addresses important gaps in current research in terms of 9 the scale of observational units (neighborhoods or zones, instead of individual households or

entire regions), while reflecting spatial autocorrelation patterns, without which the estimated

parameters may be biased. This model structure offers a solid framework for spatially-based
 choice behavior, as EV market shares rise over time, and data sets on EV ownership and
 infrastructure become more detailed.

13

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16

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- 20 21 **APPENDIX**
- 22

23 The three-level scheme relies on a series of conditional distributions, starting from the marginal

24 distribution of ϕ_3 which follows a multivariate normal distribution: $p(\phi_3) \sim N(\mu_3, \Sigma_{33})$, with

25 $\mu_3 = \mathbf{0}$ and $\mathbf{\Sigma}_{33} = [\tau_3(\mathbf{D} - \rho_3 \mathbf{W})]^{-1}$. The diagonal matrix, **D**, stores the number of neighbors

for each geographic unit along its diagonal line; τ_3 is a scaling factor to fine-tune the covariance

27 matrix, Σ_{33} ; ρ_3 measures the strength of spatial autocorrelation of the response type 3 (i.e., 28 ICEVs); and the square weight matrix, **W**, is defined by contiguity (i.e., $\mathbf{W}_{ij} = 1$ if *i* and *j* share a

border and $\mathbf{W}_{ii} = 0$ if otherwise). For ease of exposition, one can assume the following sequence

of conditional normal distributions: $p(\boldsymbol{\phi}) = p(\boldsymbol{\phi}_1 | \boldsymbol{\phi}_2, \boldsymbol{\phi}_3) \cdot p(\boldsymbol{\phi}_2 | \boldsymbol{\phi}_3) \cdot p(\boldsymbol{\phi}_3)$. Moreover,

the marginal distribution of (ϕ_2, ϕ_3) can be obtained by removing irrelevant elements (with

respect to ϕ_2 and ϕ_3) from the full distribution, leading to the following distribution:

33
$$\begin{pmatrix} \boldsymbol{\phi}_2 \\ \boldsymbol{\phi}_3 \end{pmatrix} \sim N\left(\begin{pmatrix} \boldsymbol{\mu}_2 \\ \boldsymbol{\mu}_3 \end{pmatrix}, \begin{bmatrix} \boldsymbol{\Sigma}_{22} & \boldsymbol{\Sigma}_{23} \\ \boldsymbol{\Sigma}'_{23} & \boldsymbol{\Sigma}_{33} \end{bmatrix}\right)$$

34

Conditionally, $\phi_2 | \phi_3 \sim N(\mathbf{A}_{23} \phi_3, [(\mathbf{D} - \rho_2 \mathbf{W})\tau_2]^{-1})$ (see Wang and Kockelman [2013]) where A₂₃ describes the aspatial correlation between response types 2 and 3, as well as the spatially-

37 lagged correlation between the two response types, formally: $A_{23} = \eta_{0,23}I + \eta_{1,23}W$.

38

39 Taking this conditioning one step further, $\phi_1 | \phi_2, \phi_3 \sim N(\mathbf{A}_{13}\phi_3 + \mathbf{A}_{12}\phi_2, [(\mathbf{D} - \rho_1 \mathbf{W})\tau_1]^{-1}),$

40 where A_{13} and A_{12} capture the aspatial and spatially-lagged correlation across response types 1

- 41 and 3, and response types 1 and 2, formally: $\mathbf{A}_{13} = \eta_{0,13}\mathbf{I} + \eta_{1,13}\mathbf{W}$ and $\mathbf{A}_{12} = \eta_{0,12}\mathbf{I} + \eta_{1,12}\mathbf{W}$.
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